

Data visualization

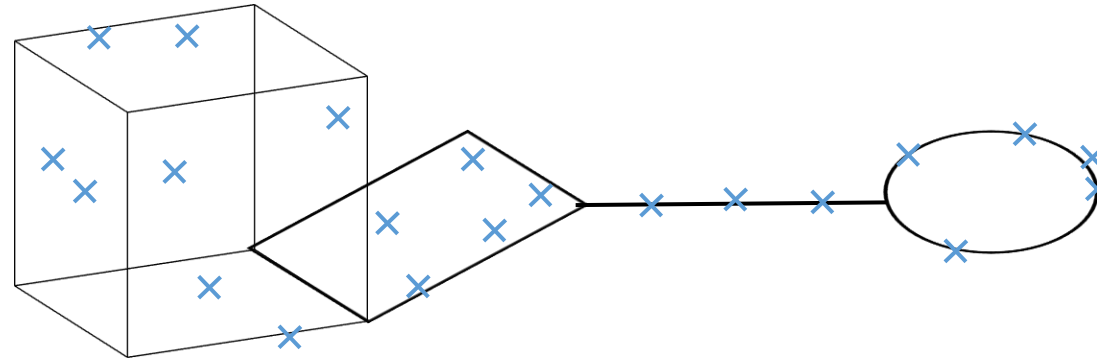
Dimension Reduction - Principal Components Analysis (PCA)

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Motivation

- **Nominal (observed)** dimensionality = number of **measurements** for each observation
- **Intrinsic (true)** dimensionality = dimension of the **space actually covered** by the observations (number of dimensions needed to describe an observation)



- **Nominal** dimensionality of a set is **higher** or equal to the **intrinsic** dimensionality → finding a **projection** from the nominal space to the intrinsic space

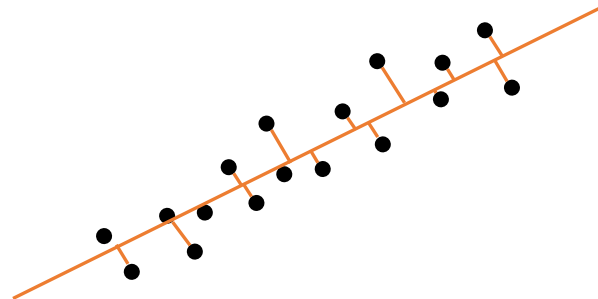
Nominal vs intrinsic dimensionality in real data

- Patients observations
 - Number of operations
 - Insurance company costs
 - Blood pressure
 - Wake-up time
 - Number of days spent in hospital



Principal components analysis

- PCA is a nonparametric **tool** for **extracting relevant information** from (usually highly dimensional data) data
- **Goal** of PCA is to find the **linear subspace** in which the data reside
 - The subspace should fit the data as best as possible
 - E.g., cloud of points along a diagonal is a linear subspace of a 2D space



Application domains

- **Machine learning**
 - Dimension reduction pre-step
- **Visualization**
 - Objects represented by many descriptors
 - PCA helps to find structure among objects which could not be visualized otherwise
- **Compression**
 - Representation of objects only by their coordinates in the respective subspace
 - E.g. in the eigenfaces (see later), each face can be reasonable approximated by only 10 coordinates

Linear algebra review

Matrices, norm, trace, eigendecomposition, spectral decomposition, SVD

Variance, covariance (1)

- **Variance** measures the spread of data in a dataset from the mean

$$\text{var}(X) = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n}$$

- **Covariance** measures how each of the dimensions varies from the mean with respect to each other

$$\text{cov}(X, Y) = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{n}$$

Variance, covariance (2)

- **Positive** covariance of two dimensions indicates that they change together (number of hours spent studying – grade)
- **Negative** covariance indicates that change in one dimension causes inverse change in the other (number of hours spent in a pub – balance of your bank account)
- **Covariance matrix** is a matrix of all pairwise covariences, e.g., for 3 dimensions X, Y, Z:

$$\begin{pmatrix} cov(X, X) & cov(X, Y) & cov(X, Z) \\ cov(Y, X) & cov(Y, Y) & cov(Y, Z) \\ cov(Z, X) & cov(Z, Y) & cov(Z, Z) \end{pmatrix}$$

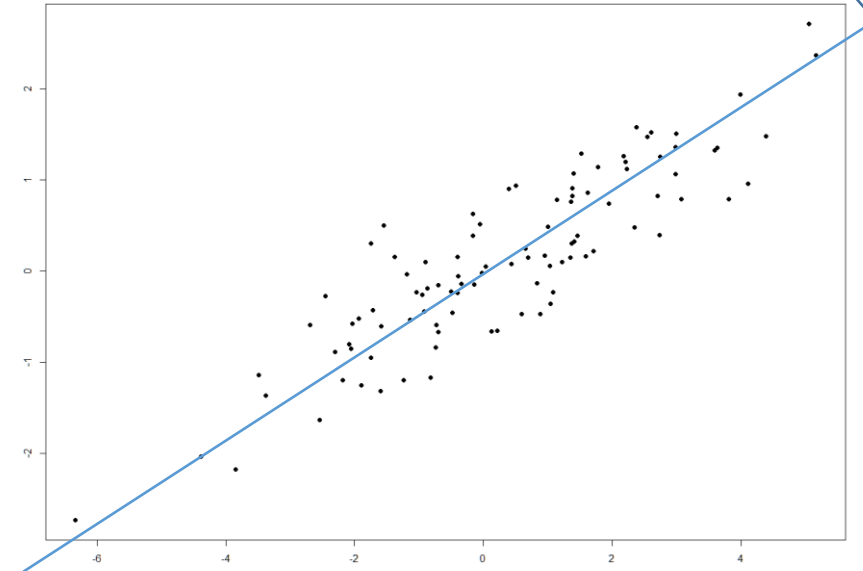
PCA formulation (1)

If we project the data onto this line, we lose as little information as possible = we keep as much variance as possible.

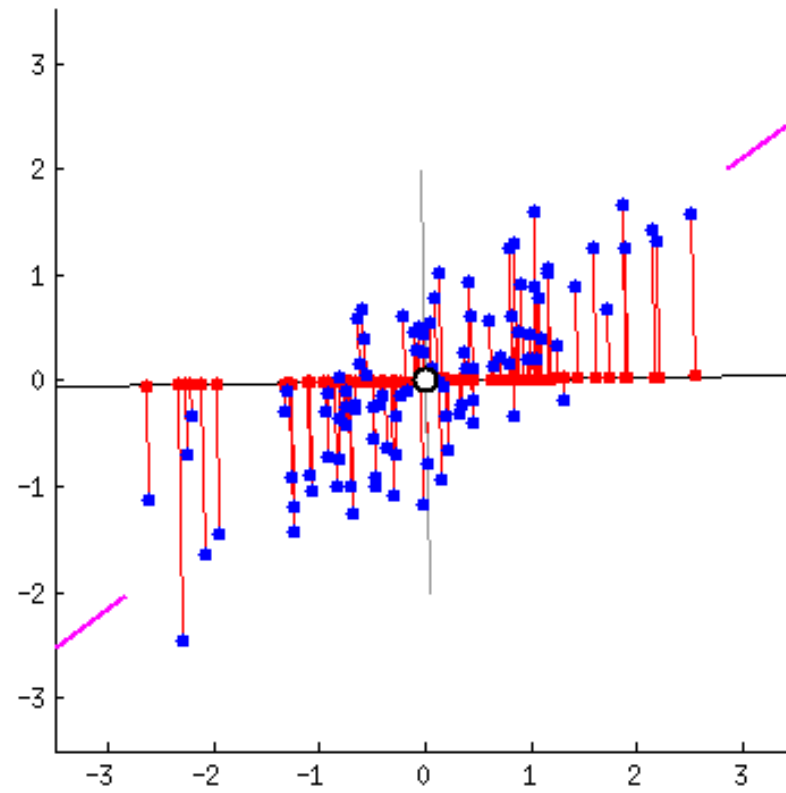
- Let us have a random variable (observations) $\mathbf{x}^T = (x_1, \dots, x_p)$ with mean $\boldsymbol{\mu}$ and covariance matrix $\boldsymbol{\Sigma}$

- First PC is the linear combination

$$y_1 = \mathbf{a}_1^T \mathbf{x} = \sum_{i=1}^p a_{1i} x_i$$



where \mathbf{a}_1 is chosen such that $\mathbf{var}(y_1)$ is maximum
subject to $\mathbf{a}_1^T \mathbf{a}_1 = \mathbf{1}$ (normalization constraint)

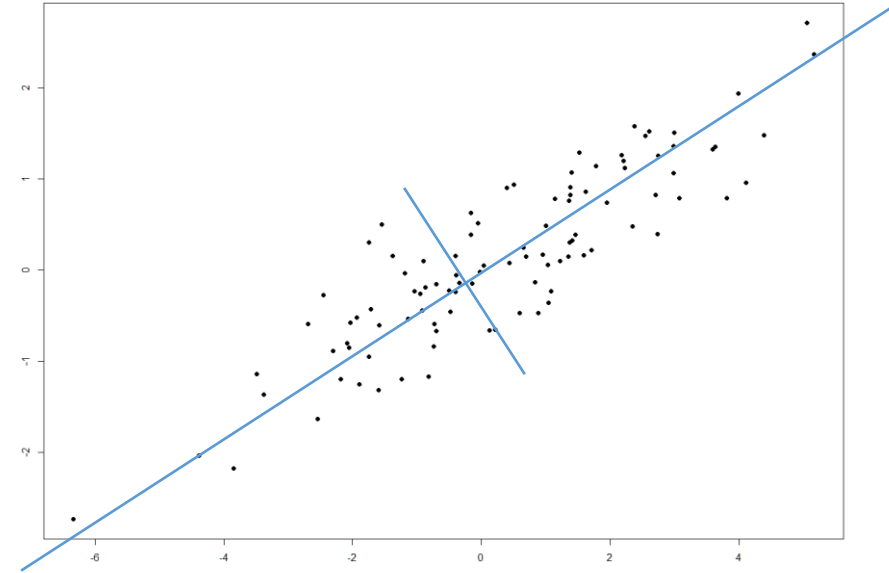


PCA formulation (2)

- Second PC is the linear combination

$$y_2 = \mathbf{a}_2^T \mathbf{x} = \sum_{i=1}^p a_{2i} x_i$$

where \mathbf{a}_k is chosen to **maximize $\text{var}(y_2)$**
subject to **$\mathbf{a}_2^T \mathbf{a}_2 = \mathbf{1}$** and **$\text{cov}(y_1, y_2) = \mathbf{0}$**



PCA formulation (3)

- Generally, k-th PC is the linear combination

$$y_k = \mathbf{a}_k^T \mathbf{x} = \sum_{i=1}^p a_k x_i$$

where \mathbf{a}_k is chosen such that $\mathbf{var}(\mathbf{y}_k)$ is maximum
subject to $\mathbf{a}_k^T \mathbf{a}_k = \mathbf{1}$ and $\forall \mathbf{l}, \mathbf{l} < \mathbf{k}: \mathbf{cov}(\mathbf{y}_k, \mathbf{y}_l) = \mathbf{0}$

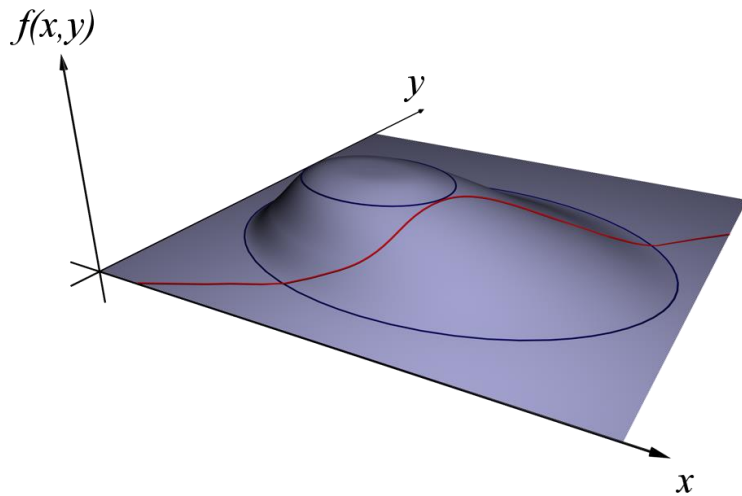
Searching for the first PC (1)

- Assumption that the data are normalized, i.e., the **mean is subtracted**
- Find a **1D subspace** so that the observations have **maximum spread** in it → maximizing variance

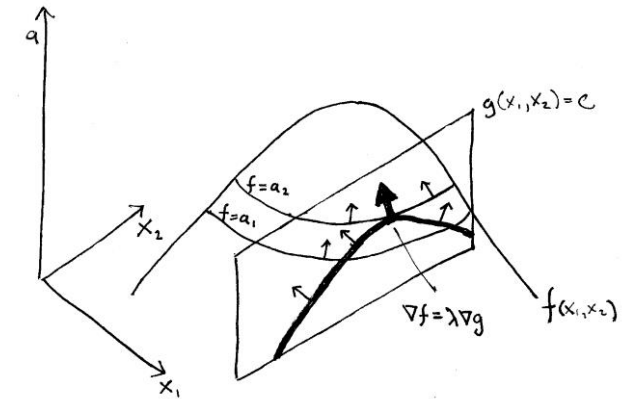
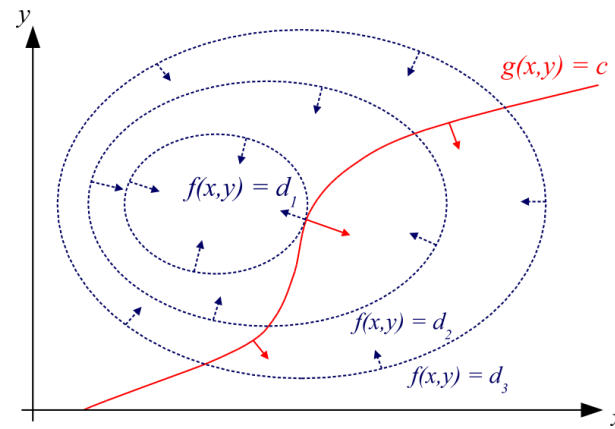
$$\begin{aligned} \mathit{var}(\mathbf{y}_1) &= \mathit{var}(\mathbf{a}_1^T \mathbf{X}) = E[(\mathbf{a}_1^T \mathbf{X} - E[\mathbf{a}_1^T \mathbf{X}])(\mathbf{a}_1^T \mathbf{X} - E[\mathbf{a}_1^T \mathbf{X}])^T] \\ &= E[(\mathbf{a}_1^T \mathbf{X})(\mathbf{a}_1^T \mathbf{X})^T] = E[\mathbf{a}_1^T \mathbf{X} \mathbf{X}^T \mathbf{a}_1] = E[\mathbf{a}_1^T \boldsymbol{\Sigma} \mathbf{a}_1] = \mathbf{a}_1^T \boldsymbol{\Sigma} \mathbf{a}_1 \end{aligned}$$

- The goal is to **maximize** variance given $\mathbf{a}_1^T \mathbf{a}_1 = 1$ → **Lagrange multipliers**

Lagrange multipliers



source: Wikipedia



source: Andrew Chamberlain (The Idea Shop)

- Maximize $f(x, y)$ subject to $g(x, y) = c \rightarrow$ introduction of a new variable - Lagrange multiplier λ ($\nabla f = \lambda \nabla g \rightarrow \nabla f - \lambda \nabla g = 0$)

Lagrangian

$$\Lambda(x, y, \lambda) = f(x, y) + \lambda(g(x, y) - c) \rightarrow \frac{\Delta \Lambda(x, y, \lambda)}{\Delta x, y, \lambda} = 0$$

Searching for the first PC (2)

- Transcription into the Lagrangian form

$$\Lambda(a_1, \lambda) = a_1^T \Sigma a_1 - \lambda(a_1^T a_1 - 1)$$

- Now we need to differentiate the Lagrangian

$$\frac{\partial \Lambda(a_1, \lambda)}{\partial a_1} = \frac{\partial \Lambda(a_1, \lambda)}{\partial \begin{bmatrix} a_{11} \\ \dots \\ a_{1k} \end{bmatrix}} = 2\Sigma a_1 - 2\lambda a_1 = 0$$

Searching for the first PC (3)

$$2\Sigma a_1 - 2\lambda a_1 = 0$$

- This leads to the eigenproblem $\Sigma a_1 = \lambda a_1 \rightarrow a_1$ is an eigenvector of Σ with eigenvalue λ

$$\mathit{var}(\mathbf{y}_1) = \mathit{var}(\mathbf{a}_1^T \mathbf{X}) = \mathbf{a}_1^T \Sigma \mathbf{a}_1 = \lambda \mathbf{a}_1^T \mathbf{a}_1 = \lambda$$

- Suppose that $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \rightarrow$ to maximize $\mathit{var}(\mathbf{y}_1)$ we must choose $\lambda = \lambda_1$

Searching for the next PCs

- The principle is similar, but due to the uncorrelation requirement we must extend the constraint with

$$0 = \text{cov}(y_1, y_2) = \text{cov}(a_1^T x, a_2^T x) = a_1^T \Sigma a_2 = a_2^T \Sigma a_1 = a_2^T \lambda a_1 = \lambda a_2^T a_1$$

- Leading to a modified Lagrangian

$$\Lambda(a_2, \lambda, \kappa) = a_2^T \Sigma a_2 - \lambda(a_2^T a_2 - 1) - \kappa(a_2^T a_1)$$
$$\left(a_2^T \Sigma a_2 - \lambda(a_2^T a_2 - 1) - \kappa(a_2^T a_1) \right) \frac{d}{da_2} = 0$$

$$\begin{aligned} \Sigma a_2 - \lambda a_2 - \kappa a_1 &= 0 \\ a_1^T \Sigma a_2 - \lambda a_1^T a_2 - \kappa a_1^T a_1 &= 0 \\ 0 - 0 - \kappa &= 0 \end{aligned}$$

$$\begin{aligned} \Sigma a_2 - \lambda a_2 &= 0 \\ \Sigma a_2 = \lambda a_2 &\Rightarrow \lambda = \lambda_2 \end{aligned}$$

PCA transformation

- Thus, the **coefficients of the linear combination which transform the observations onto the PCs are formed by eigenvalues of the covariance matrix**
- Let A contain the eigenvectors a_i as its columns and let x be a p -dimensional vector representing an observation, then

$$y = A^T(x - \mu)$$

Variance

- **PCs are components of variance explaining** the total variation in the data

- The sum of variances of the original variables $var(X)$ and of the PCs $var(Y) = var(AX)$ are the same

$$\Sigma = A\Lambda A^T$$
$$tr(\Sigma) = tr(A\Lambda A^T) = tr(\Lambda A^T A) = tr(\Lambda)$$

- Therefore

$$\frac{\lambda_i}{\lambda_1 + \dots + \lambda_p}$$

can be interpreted as the total variation in the original data explained by the i -th principal component

Scores and loadings

- **Scores**

- **Transformed variable values** corresponding to a particular observation
 - Original data multiplied by the loadings
- Geometrically, scores are the **coordinates** of each observation with respect to the **new axis**

- **Loadings**

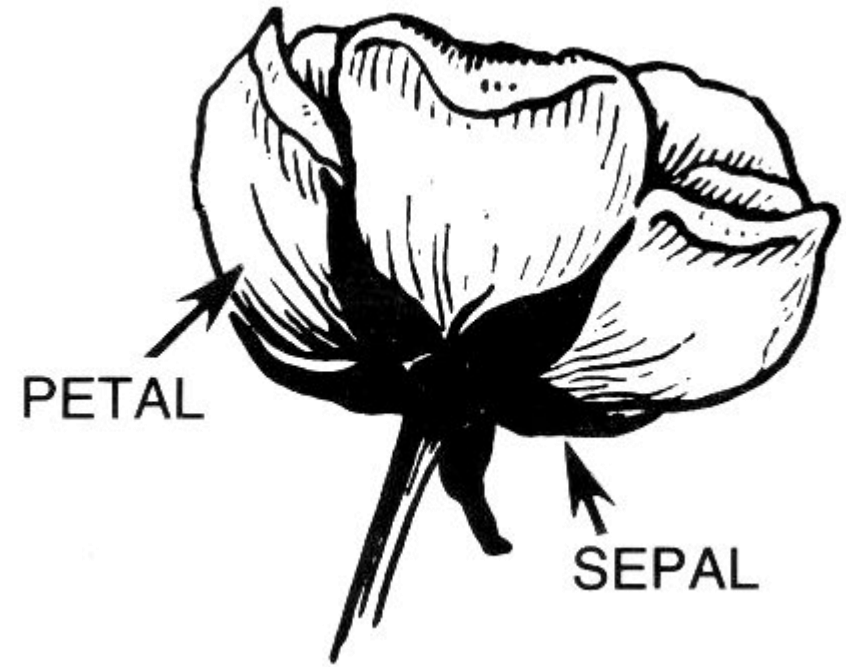
- Weight by which each standardized original variable should be multiplied to get the component score → separate loadings for each component
- Expresses which variables have high **loading** in which PCs
 - Loadings close to zero indicate which variables do not contribute much to given component
- Extent to which given **variable** is **correlated with given component**

Scale invariance

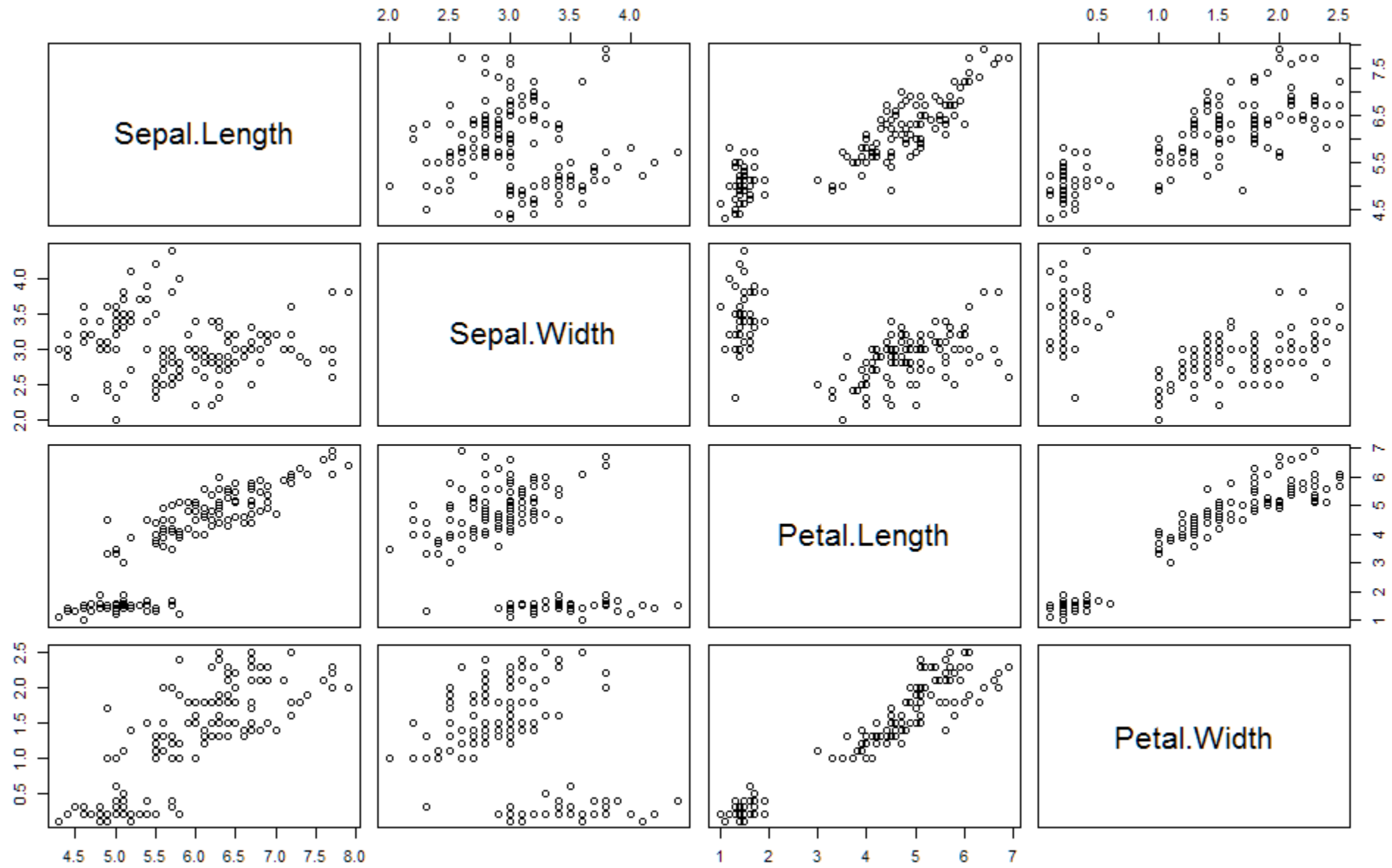
- PCA is **NOT** scale invariant → variance in consistently large variable will dominate the spectrum of eigenvalues → variables should be of comparable scale
 - E.g., if height of a person was expressed in nanometers, the first PC would probably be identical with the height dimensions (highest variance)
- Often the variables are divided by the square root of its variance → **correlation matrix** instead of covariance matrix

$$\text{cor}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{cov}(X, X)\text{cov}(Y, Y)}}$$

Iris dataset



- One of the [R datasets](#)
- The measurements in centimeters of the variables sepal length and width and petal length and width, respectively, for 50 flowers from each of 3 species of iris $\rightarrow n = 150, p = 4$



PCA in R

- The most common ways to conduct PCA in R is `prcomp` (stats), `princomp` (stats) or `PCA` (FactoMineR)

```
data(iris)
```

```
ir.descriptors <- iris[, 1:4]
```

```
ir.species <- iris[, 5]
```

```
ir.pca <- prcomp(ir.descriptors, center = TRUE, scale. = TRUE)
```

```
print(ir.pca)
```

```
summary(ir.pca)
```

Standard deviations:

```
[1] 1.7083611 0.9560494 0.3830886 0.1439265
```

Rotation:

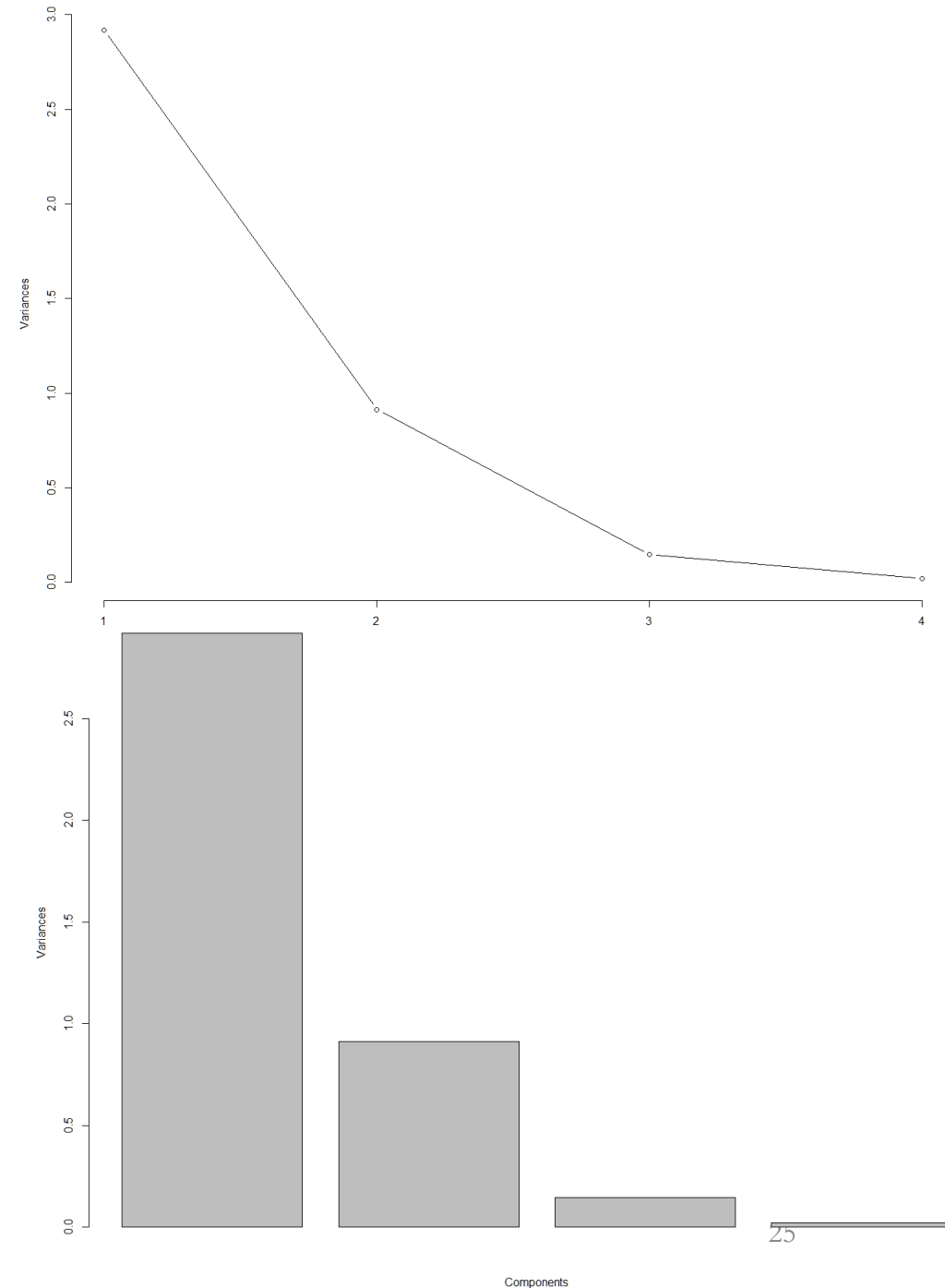
	PC1	PC2	PC3	PC4
Sepal.Length	0.5210659	-0.37741762	0.7195664	0.2612863
Sepal.Width	-0.2693474	-0.92329566	-0.2443818	-0.1235096
Petal.Length	0.5804131	-0.02449161	-0.1421264	-0.8014492
Petal.Width	0.5648565	-0.06694199	-0.6342727	0.5235971

Importance of components:

	PC1	PC2	PC3	PC4
Standard deviation	1.7084	0.9560	0.38309	0.14393
Proportion of Variance	0.7296	0.2285	0.03669	0.00518
Cumulative Proportion	0.7296	0.9581	0.99482	1.00000

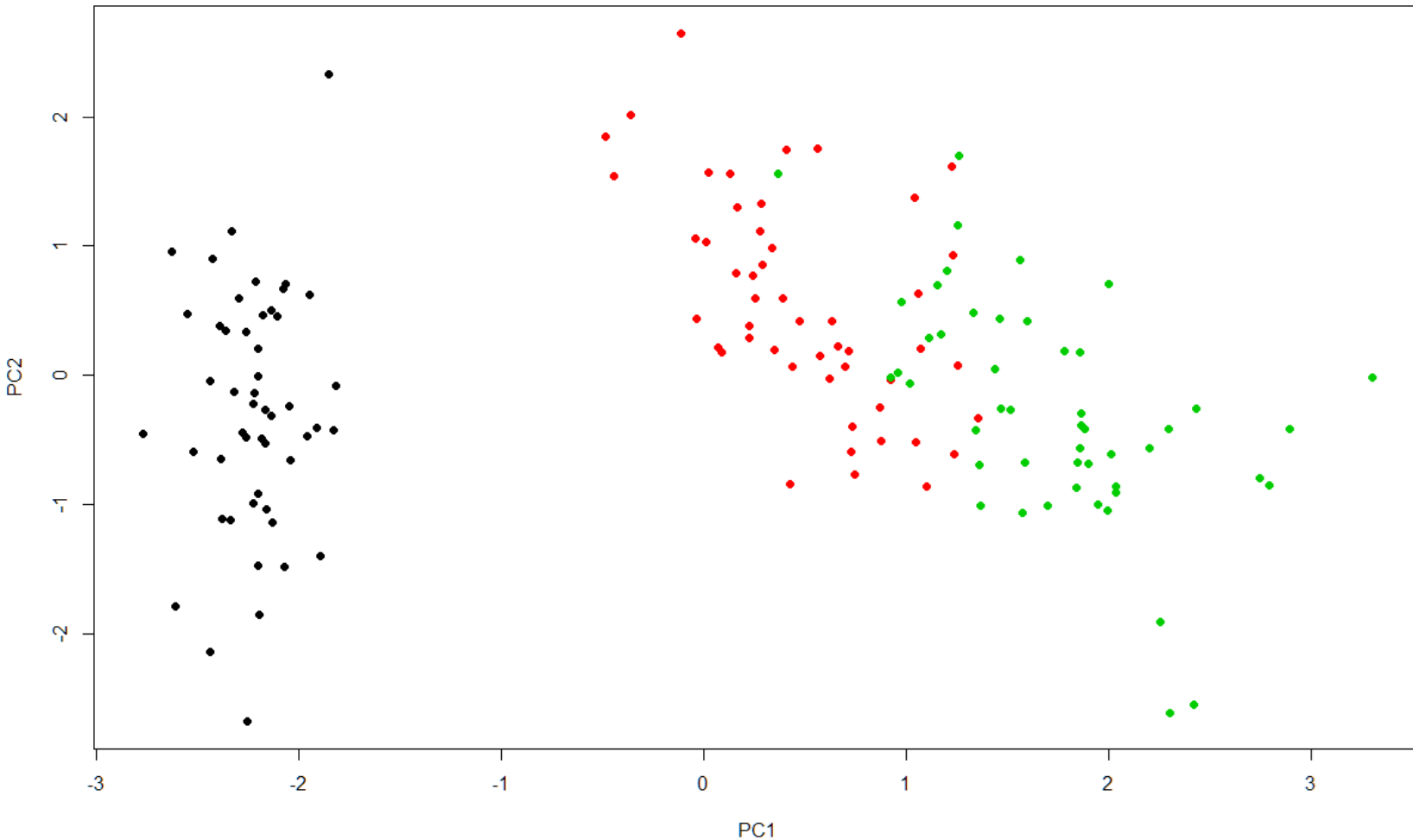
Scree plot

- Display of variance of each of the component
- Plot of magnitudes of eigenvalues
- Gives impression of the intrinsic dimensionality



Score plot

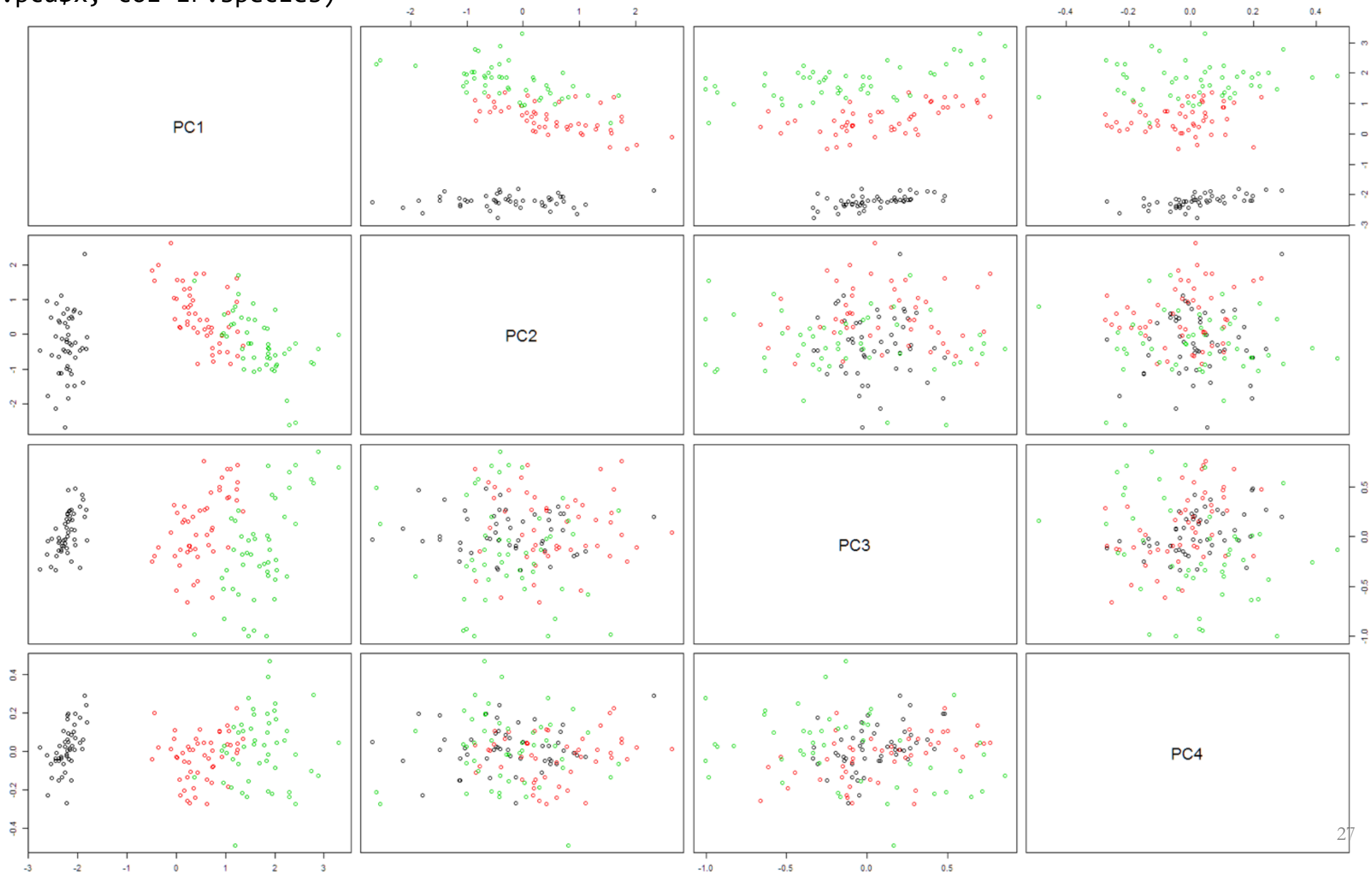
- Closeness in the score plot indicates similar “behavior” between samples



```
ir.pca$x
```

```
plot(ir.pca$x, col=ir.species,  
     pch = c(16))
```

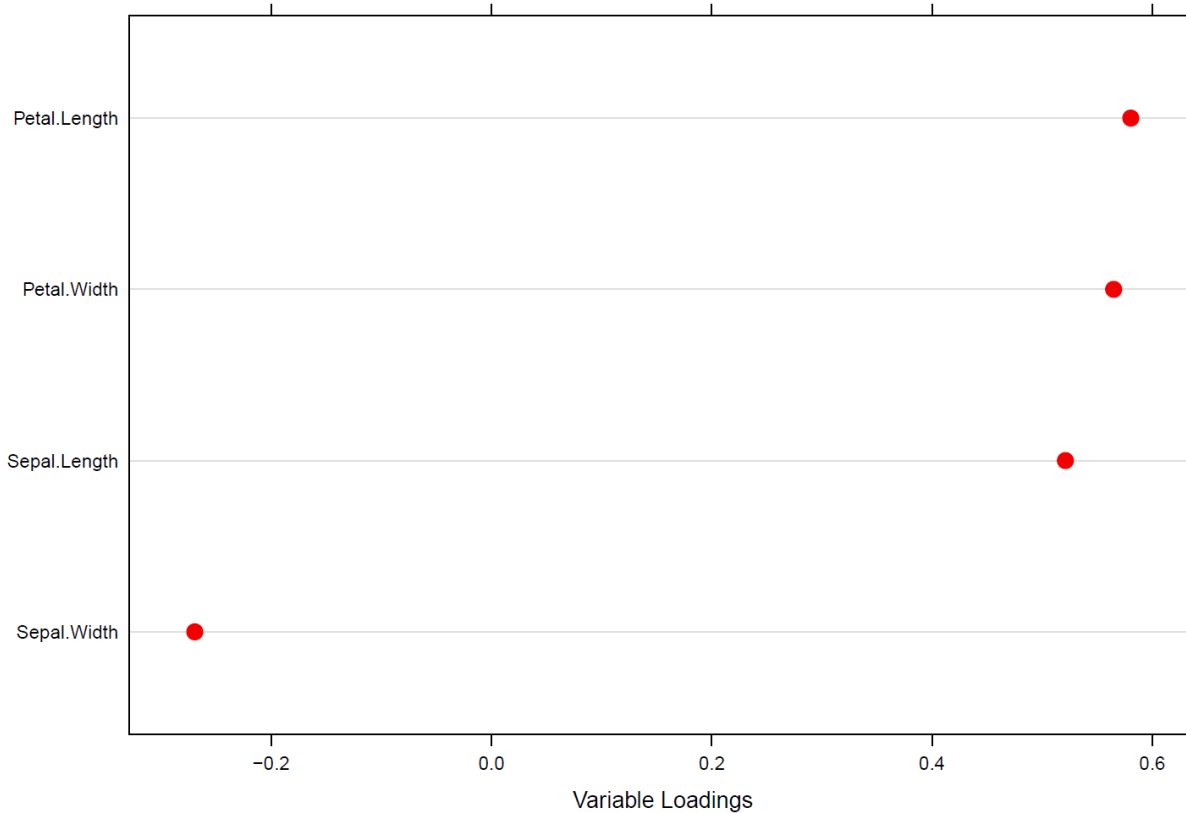
```
pairs(ir.pca$x, col=ir.species)
```



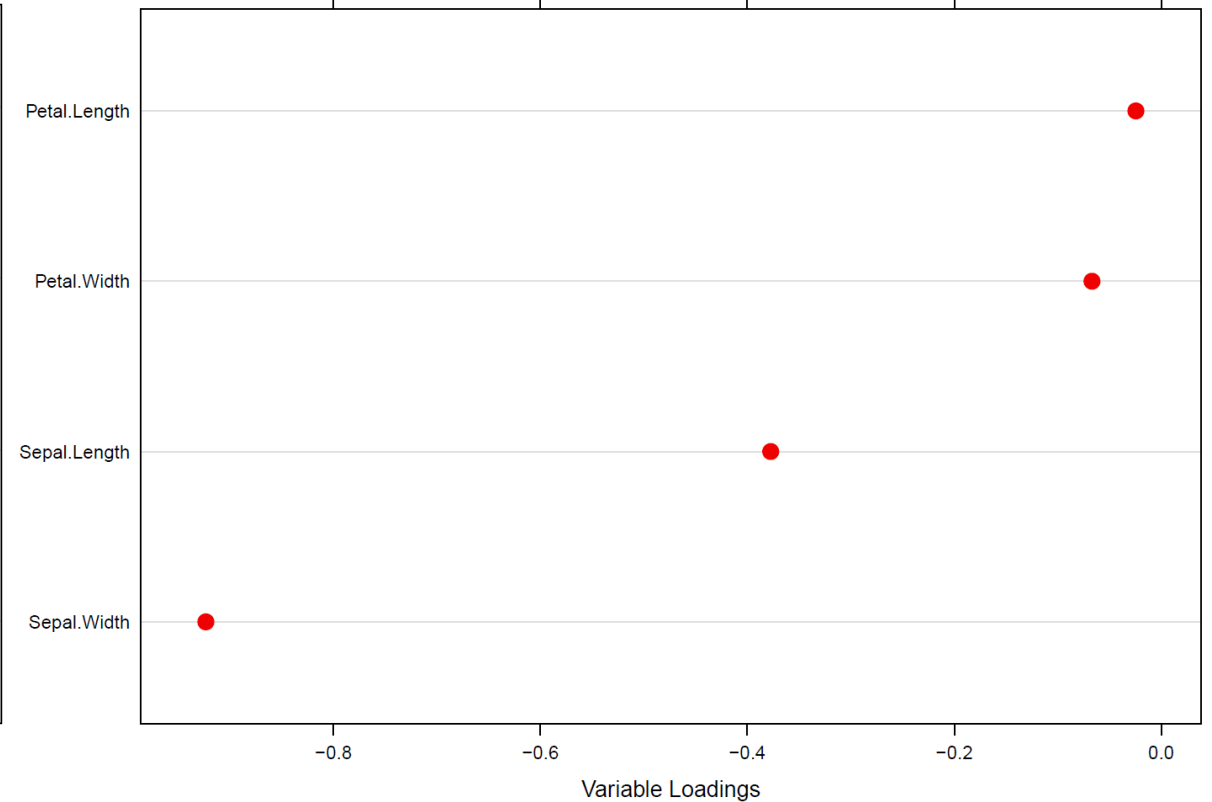
Loadings plot

- Closeness in the score plot indicates similar “behavior” between variables

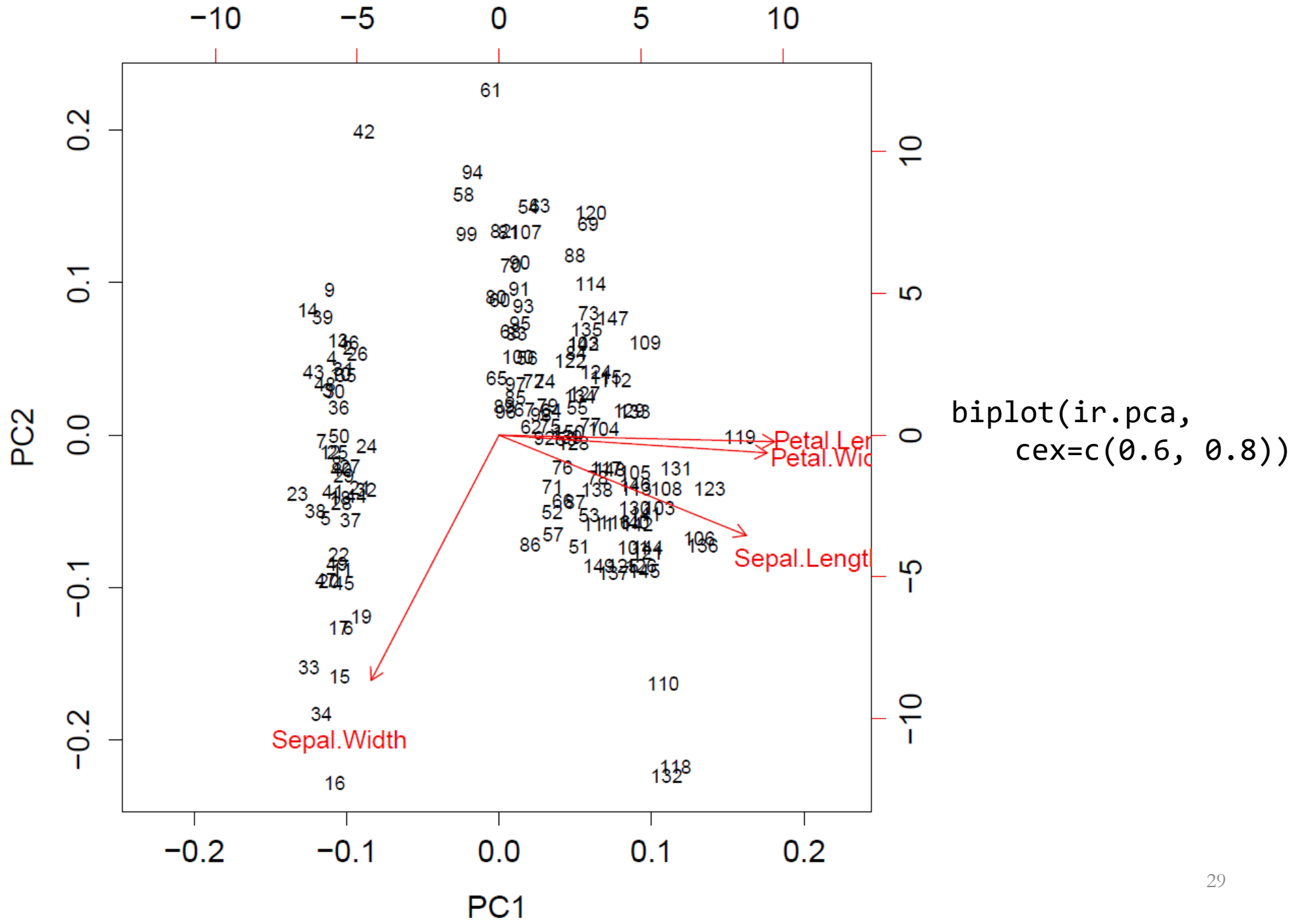
Loadings Plot for PC1



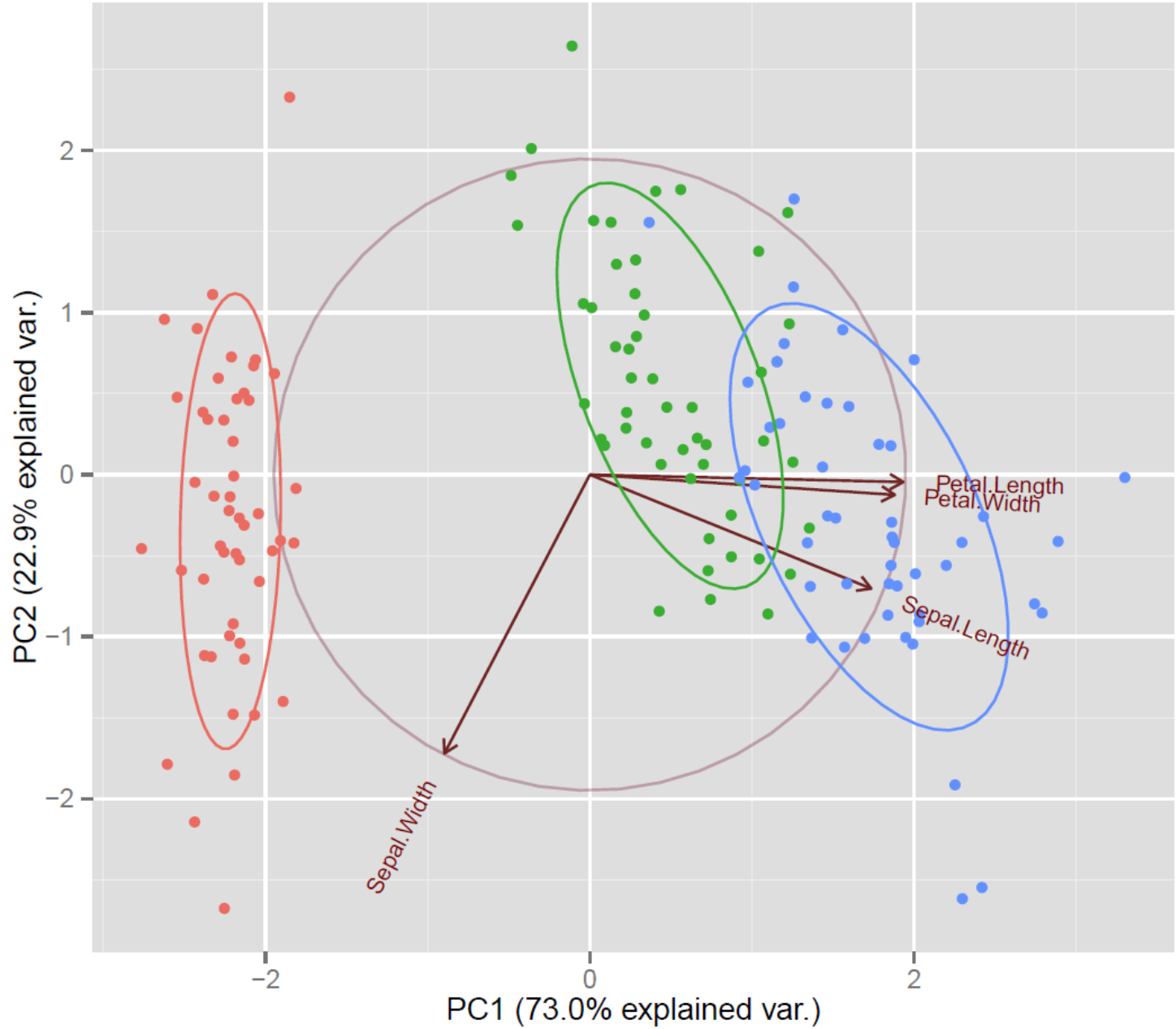
Loadings Plot for PC2

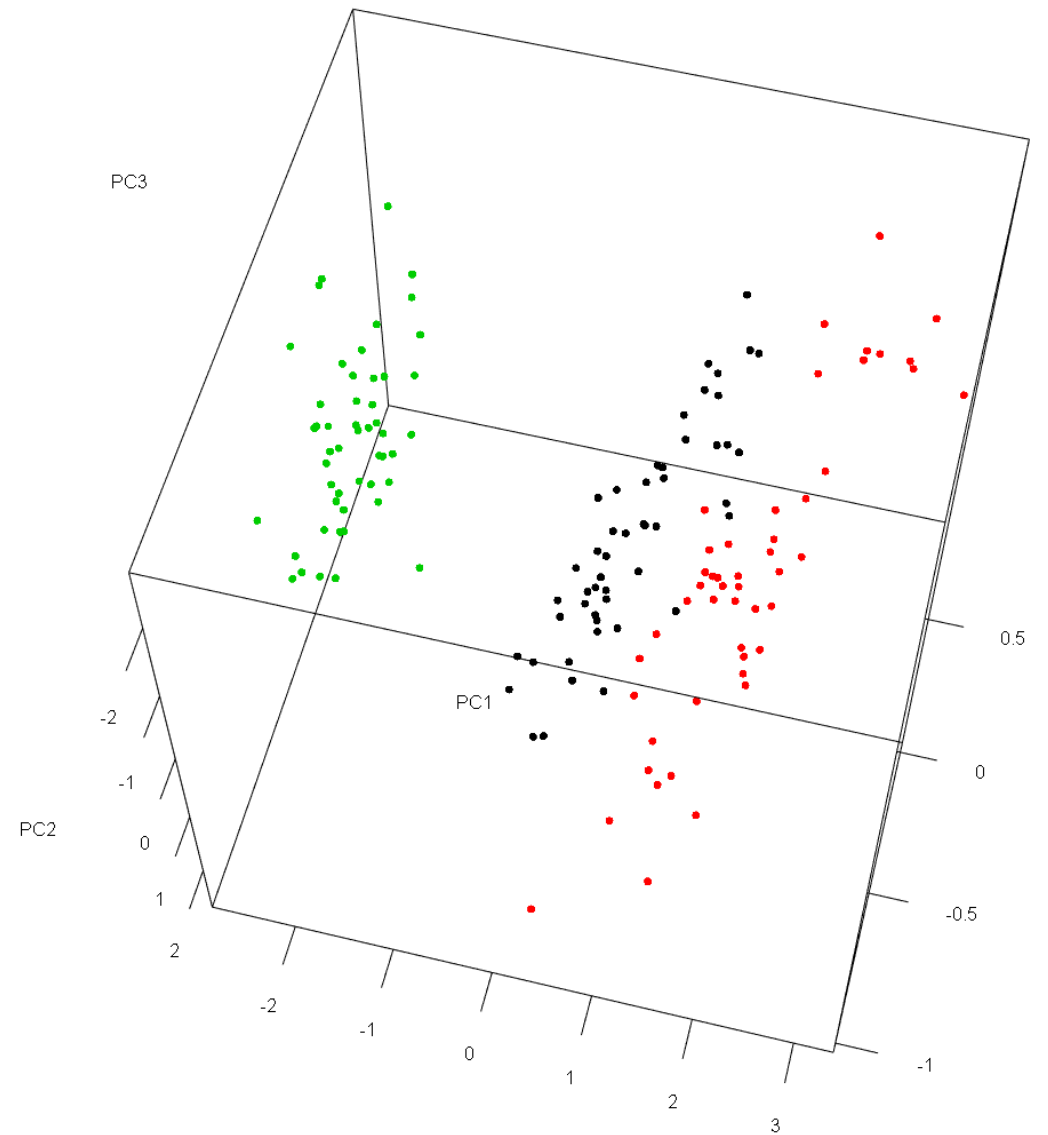
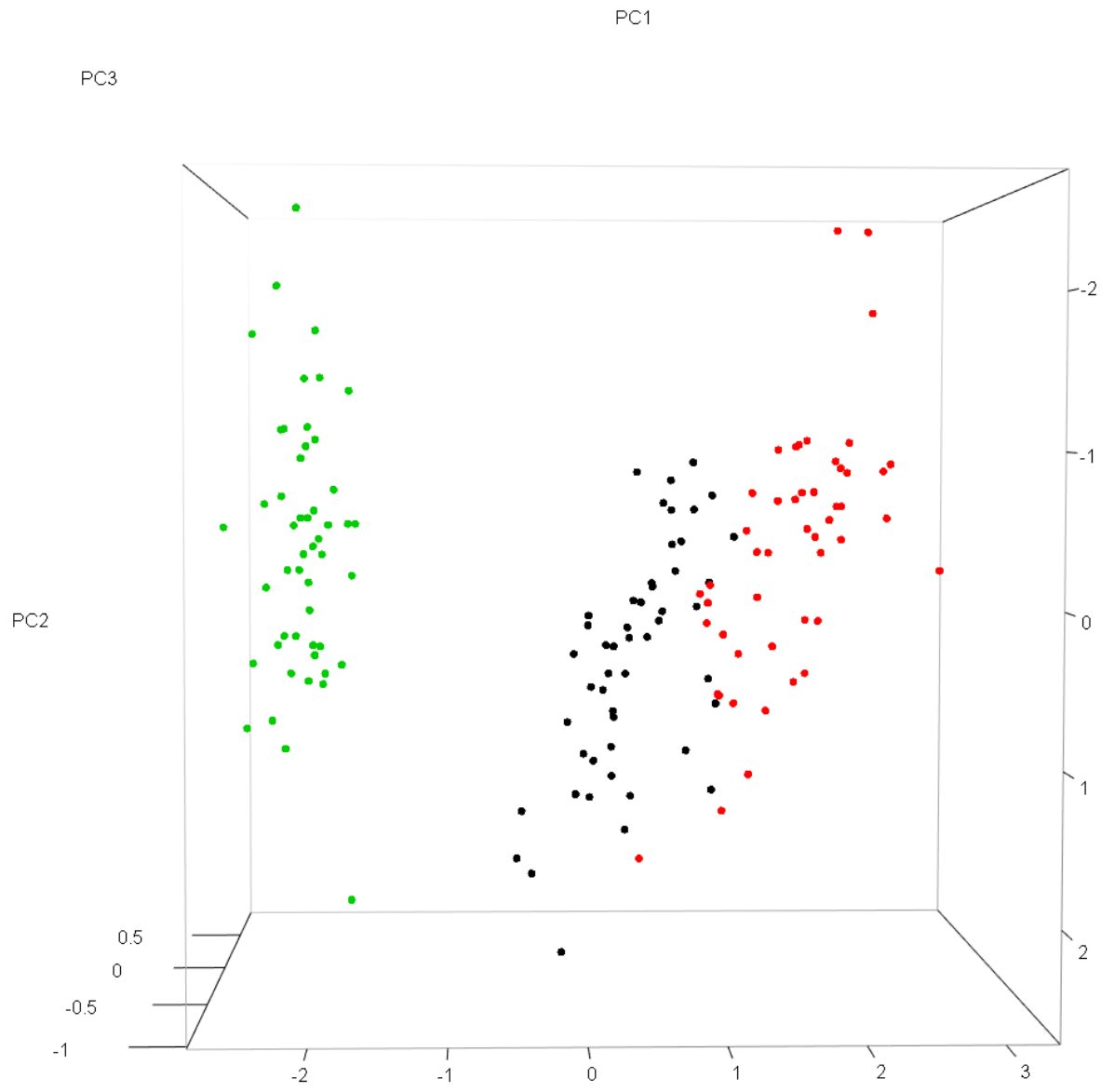


Biplot



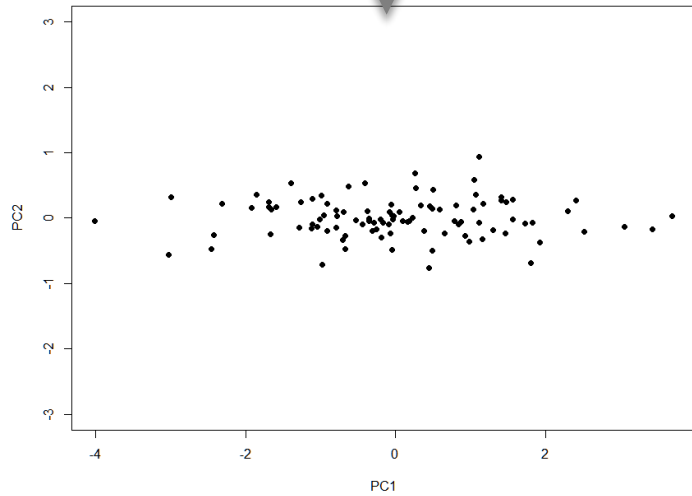
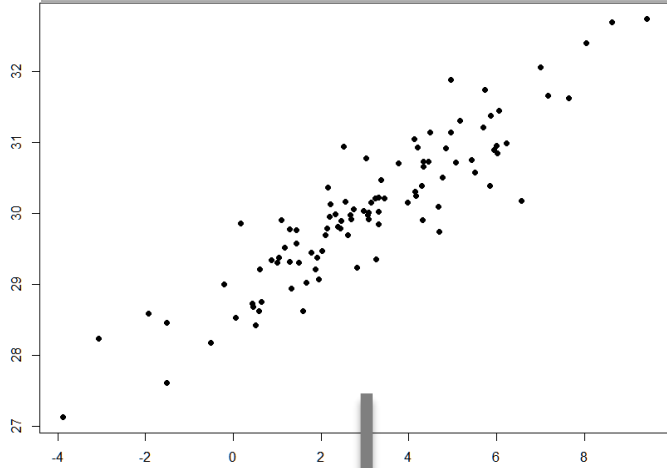
—●— setosa —●— versicolor —●— virginica





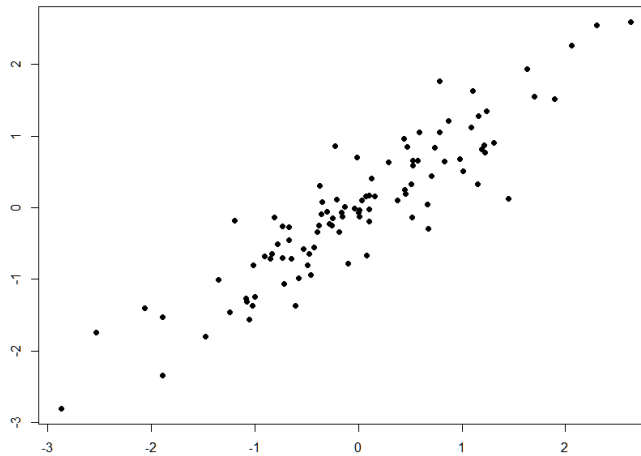
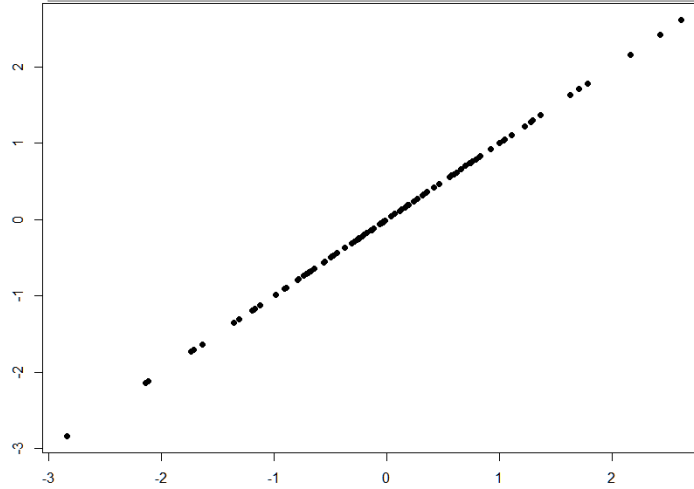
```
plot3d(ir.pca$x[,1:3], col=as.numeric(factor(iris$Species, levels =
c("versicolor", "virginica", "setosa"))), size=7)
```

```
data.orig = rmvnorm(100, mean =  
c(3, 30), sigma = matrix(c(5, 2,  
2, 1), nrow = 2))  
plot(data.orig, pch=c(16))
```



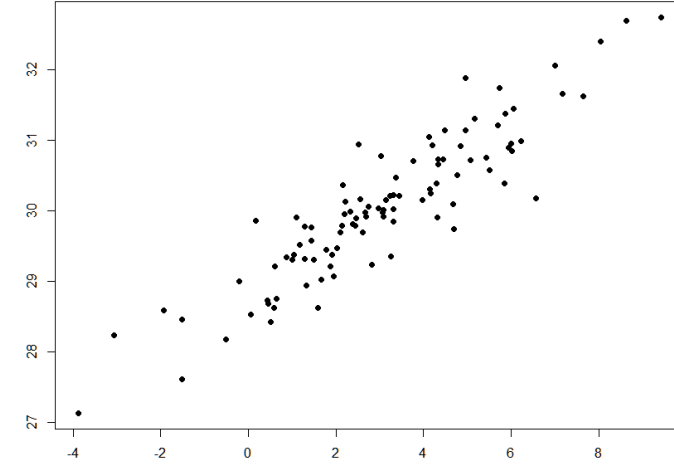
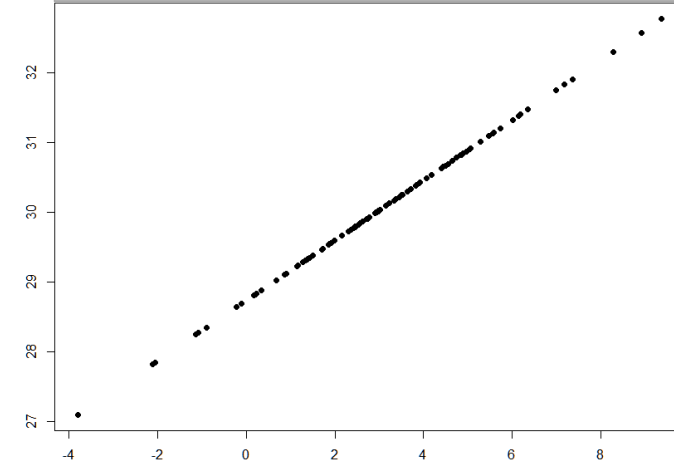
```
data.pca <- prcomp(data.orig,  
center = TRUE, scale = TRUE)  
plot(data.pca$x, pch=c(16),  
ylim=c(-3,3))
```

```
restr = data.pca$x[,1] %**  
t(data.pca$rotation[,1])  
plot(restr, pch=16)
```



```
restr = data.pca$x %**  
t(data.pca$rotation)  
plot(restr, pch=16)
```

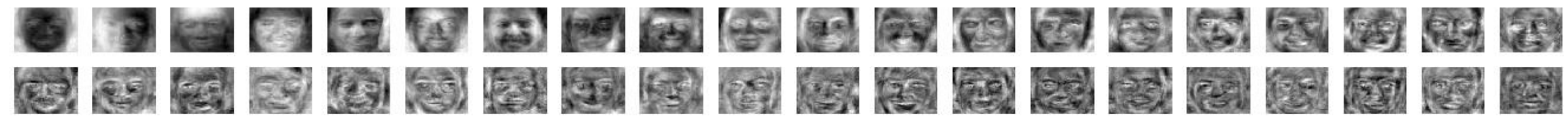
```
restr <- scale(restr, center =  
FALSE , scale=1/data.pca$scale)  
restr <- scale(restr, center = -1  
* data.pca$center, scale=FALSE)  
plot(restr, pch=16)
```



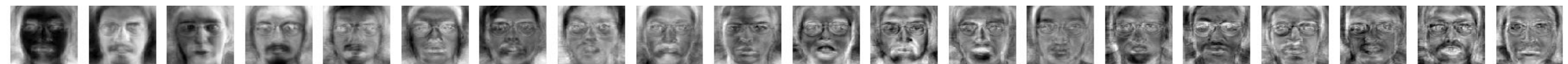
```
restr <- scale(restr, center =  
FALSE , scale=1/data.pca$scale)  
restr <- scale(restr, center = -1  
* data.pca$center, scale=FALSE)  
plot(restr, pch=16)
```


PCA on grayscale images

- [Dataset](#) of 96x96 grayscale images
- PCA allows to compress images by representing the original pixels by few linear combinations (scores)
 1. Convert each image into a 9216-long (96x96) vector of numbers (0-255) → each image is a point in a 9216-dimensional space
 2. Run PCA on the 9216-dimensional objects
 3. Take first k PCs (first k columns of the matrix $A \rightarrow A_k$) so that enough variability is captured
 4. Convert each object x into the new k -dimensional space using $A_k x$



Principal components



Approximations



Literature

- Jolliffe, I.T. (2002) *Principal Component Analysis, Second Edition.* Springer-Verlag New York